


Comparison of self-reported and directly measured weight and height among women of reproductive age: a systematic review and meta-analysis

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Key words

Body height, body mass index, bodyweight, bodyweights and measures, reproductive age, self-assessment, women

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Conflict of interest

The authors have stated explicitly that there are no conflicts of interest in connection with this article.

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Introduction

Body mass index (BMI) is a simple and useful indicator to classify individuals as healthy or at risk according to their weight and height (1). Traditional anthropometric

Abstract

Introduction. The use of self-report as a strategy for collecting data on women's weight and height is widespread in both clinical practice and epidemiological studies. This study aimed to compare self-reported and directly measured weight and height among women of reproductive age. **Material and methods.** In July 2015 we searched MEDLINE, EMBASE, COCHRANE, CINHAL, LILACS and gray literature. We included women of reproductive age (12–49 years old) independently of their weight or height at the time of the study. Women with any condition that implies regular tracking of their weight (for example, eating disorder) were excluded. Two reviewers independently selected, extracted and assessed the risk of bias of the studies. We used REVMAN 5.3 to perform the meta-analysis. Heterogeneity was assessed using the I^2 statistic. **Results.** Following eligibility assessment, 21 studies of 18 749 women met the inclusion criteria. The results of the meta-analysis showed an underestimation of weight by -0.94 kg (95% CI -1.17 to -0.71 kg; $p < 0.0001$; $I^2 = 0\%$) in the overall sample and an overestimation of height by 0.36 cm (95% CI 0.20 – 0.51 ; $p < 0.0001$; $I^2 = 35\%$) based on self-reported vs. directly measured values. **Conclusion.** This review shows that self-reported weight and height of women of reproductive age differs slightly from direct measures. We consider that the magnitude at which self-reported data over- or underestimates the real value, is negligible regarding clinical and research use.

Abbreviations: BMI, body mass index; CI, confidence interval; SE, standard error.

Key messages

Self-reported weight and height in women of reproductive age is a measure that closely estimates the real values and can be used as proxy both in clinical and research evaluations related to reproductive health.

measures such as weight and BMI are often used in epidemiological studies to assess changes in population health and nutritional status (2). Regarding women's health, BMI prior to pregnancy requires strict attention as it can be a risk factor not only for women but also for future generations (3). Because of this, the International Federation of Gynecology and Obstetrics (FIGO) emphasizes the need to control preconceptional body-weight and BMI to prevent abnormal values that can impact significantly on maternal and neonatal health outcomes (3).

Anthropometric measures are often gathered through self-administered questionnaires. This data collection method has the advantages of being quick, easy to administer and cost-effective when working with large samples or when individuals are spread over large areas (4). In research, the self-report of height or weight is widely used in descriptive studies to save a significant amount of time and resources (5–8). In clinical practice, self-reported measures of weight are also a useful strategy to determine historical weights; for example, self-report allows for estimation of pregnancy weight gain that would otherwise be difficult due to the variable stages at which the first antenatal visit occurs. Despite these advantages, the utility of self-reported measures has been questioned, particularly when as related to anthropometric measures. There is a global preconceived idea that participants tend to overestimate their height and underestimate their weight, resulting in a lower estimate of BMI (4). The greatest hazard of unreliable reporting of weight and height is the inaccurate estimation of the prevalence of overweight and obesity, which can result in unsupported decision-making (4).

It is vital to have an up-to-date systematic review on this topic to reduce the risk of bias when reporting the results of a study. Any important difference between self-reported and directly measured data should be taken into consideration when selecting data collection methods for future studies or clinical actions.

The objective of this review is to compare self-reported with directly measured weight and height among women of childbearing age. The purpose of these meta-analyses is to give a summary estimate of the possible bias that can occur when using self-report as a data collection method.

Material and methods

We followed the Preferred Reporting Items for Systematic Reviews and Meta-Analyses Statement (PRISMA Statement) (9,10) and the Meta-analysis of Observational Studies in Epidemiology (MOOSE) statement.

Criteria for considering studies for this review

We selected cross-sectional and prospective or historical cohort studies that compared individual self-reported with directly measured weight and height data. We included published or unpublished studies from 2000 onward that reported at least 20-paired values of self-reported and directly measured weight or height, or data of the difference between them. No language restriction was used.

We included healthy non-pregnant women of reproductive age, independent of their weight or height. We considered reproductive age to be from 12 to 49 years old. All methods to obtain a self-reported or directly measured weight and height were accepted. We excluded women with a disease or condition that implied regular monitoring or records of their weight, such as women following dietary plans or women with eating disorders.

Studies were included only if they expressed the outcome as “mean self-reported weight or height”, “mean directly measured weight or height” or “mean difference between self-reported and directly measured weight or height”.

Search methods for identification of studies: electronic searches

A literature search for articles published from 1 January 2000 to 14 July 2015 was conducted within the main international and regional databases, through generic and academic internet searches and through meta-search engines.

We searched records from the following databases:

- CENTRAL: The Cochrane Library (last available Issue 2015)
- MEDLINE (January 2000 to July 2015)
- EMBASE (January 2000 to July 2015)
- LILACS: Latin American and Caribbean Health Science Literature (January 2000 to July 2015).
- CINAHL (January 2000 to July 2015)

The simplified and complete search with filters in MEDLINE is described below; these were adapted appropriately for each database (Appendix S1). We also reviewed the reference lists of included studies for potential additional studies.

Data collection and analysis

Selection of studies. All phases of the study selection and processing were completed using EROS[®] (Early Review Organizing Software, IECS, Buenos Aires), a web-based platform designed for the process of systematic reviews (11). As an initial screening, a pair of reviewers

(M.S., N.M.) independently reviewed the articles, evaluating the titles and abstracts of identified studies according to prespecified criteria. Discrepancies were resolved by consensus of the whole research team. Articles included after the initial evaluation were retrieved in full text for a second screening to determine eligibility. Finally, the same reviewers independently extracted and assessed the risk of bias of each full-text article.

Data extraction and management. We used a web-based spreadsheet to extract the information. One reviewer extracted data from the included studies and a second reviewer double-checked it to minimize potential errors. This process was piloted on 20 papers to refine it. Discrepancies were resolved by consensus of the whole team.

The information extracted from each study included author, publication year, type of study, region and country of study, participant characteristics (age and education level), sample size, methods to obtain directly measured weight and height (stadiometer, anthropometer or other type of measuring (tape or ruler, variety of scales), methods to obtain self-reported weight and height (long distance survey, on-site interview, self-administered questionnaire), time between collection of self-reported and directly measured data, order of measures, ethical considerations and outcomes [mean self-reported and directly measured weight or height or mean differences between self-reported and directly measured weight or height, and their standard deviation (SD)]. Authors of studies reporting incomplete information were contacted to provide missing information. We waited for one month for the author's answer before excluding the article.

Assessment of risk of bias and data analysis. The risk of bias of observational studies was assessed using a checklist of essential items based on the STROBE (Strengthening the Reporting of Observational Studies in Epidemiology) (12). The STROBE essential checklist includes: selection of participants, control of confounders, measurement of exposure and outcome, and conflict of interest. Pairs of independent reviewers assessed the methodological quality. Discrepancies were resolved by consensus of the whole team.

The null hypothesis when comparing self-reported and directly measured weight and height stated no difference between methods (self-reported = directly measured). Those measurements expressed in pounds or inches were transformed to kilograms and centimeters, respectively, and the reported standard errors (SE) were converted to standard deviations using the following formula: $\sqrt{n} \times SE$. We performed a meta-analysis using the

continuous outcomes of all the studies that reported mean values of weight or height using self-reported and directly measured methods. A summary estimate obtained from the meta-analysis of a mean difference not equal to 0 would indicate that the use of self-report affects positively or negatively the measure compared with the use of direct measurements; based on either difference, self-reported values could be defined as a weak method for data collection. We used REVMAN 5.3 (13) to perform the meta-analysis and to calculate the two-tailed p -values and 95% confidence intervals (CI).

We measured heterogeneity using the I^2 statistic as follows: low heterogeneity ($I^2 < 25\%$), moderate heterogeneity ($I^2 = 25\text{--}75\%$), and high heterogeneity ($I^2 > 75\%$).

For those studies that only reported mean differences between methods, we performed a generic inverse-variance meta-analysis, which considered mean difference and SE. To be able to include all the studies, we used REVMAN's calculator function to extract mean differences and SE for each of them. The resulting value indicated the directionality of the findings. A result under 1 indicated that the directly measured values were higher than the self-reported ones; a result above 1 indicated that the self-reported values were higher than the direct measured ones; and a value of 1 indicated no difference between methods.

Prespecified subgroup analyses by age, time between self-reported and direct measurements (same day, different days), region of the study (Latin America & Caribbean, Europe, North America, Oceania, Asia), self-report method (long-distance survey, self-administered questionnaire on-site, in-person interview) and women's BMI were performed. For all the meta-analyses we used a random effect model to address possible clinical or methodological heterogeneity between studies.

We compiled the age data into three groups: 12–18, 19–35 and 36–49 years. For studies in which age was grouped differently and data could not be disaggregated, we based our groups on the category to which the majority of study participants belonged. BMI was classified following World Health Organization categories (underweight <18.5 , normal weight 18.5 to <25 , overweight 25 to <30 , and obesity ≥ 30) (1). The protocol was registered in PROSPERO, an international prospective register of systematic review protocols (Registration Number CRD42015029142).

Results

Description of studies

Results of the search. The search strategy retrieved 1638 references after removing duplicates. Of those, 1476

references were excluded by title and abstract, leaving 162. Two full texts were not found (14,15) and 139 studies did not meet the inclusion criteria. After assessment, 21 studies with 18 749 women were included in the review (Figure 1) (6,16–35).

Included studies. Of the 21 included studies, six were from Latin America and the Caribbean ($n = 3470$, 14.8% of the women) (18,29,30,32,34,35), nine from Europe ($n = 8459$, 36.2% of the women) (16,19–22,24,26,28,33) and four from North America ($n = 8264$, 35.3% of the women) (6,17,23,25). We only included one article from Oceania (31) and one from Asia (27) ($n = 3206$, 13.7% of the women). Two of the included studies were prospective cohorts (19,33) and the rest ($n = 19$) were cross-sectional studies (Table 1) (6,16–18,20–32,34,35).

Eighteen studies reported details of the tools used for self-reported and directly measured weight and height of participants (6,16–22,24–29,31,33–35). For directly measured data, height was most commonly measured by stadiometer, anthropometer or some type of measuring tape or ruler with an error of 0.1–0.5 cm; weight was measured by a variety of scales with an error of 0.1 kg (balance beam, digital or portable). Twelve of the 21 studies used self-administered on-site questionnaires as the self-reported method (6,17,20–22,24–27,33–35). Three studies gathered information via an online survey or telephone (18,19,31) and three other studies performed an in-person interview

to obtain these data (16,28,29). The remaining three studies did not report the type of method used (23,30,32). All the studies obtained the self-reported value prior to the directly measured data (6,16–35).

From the included studies, 19 reported the mean value of self-reported and directly measured weight and height (6,16–21,23–28,30–35). Two studies only reported the mean difference between methods, calculated as self-reported minus directly measured values (22,29). Only two studies showed data by BMI categories (28,29).

Risk of bias assessment. The risk of bias assessment found six studies with high risk of bias in the selection of participants (29.0%) (24,25,28,30,33,35) and two studies with the control of confounders (9.50%) (Table S1) (30,35).

Weight. According to the meta-analysis, we found that in the overall sample, the mean difference between self-reported and direct measured data for women's weight was -0.94 kg (19 studies; 16 578 participants; 95% CI -1.17 to -0.71 kg; $p < 0.0001$; $I^2 = 0\%$) (6,16–21, 23–28,30–35). When analyzed by age subgroups, we found that self-reported weight was lower than directly measured weight in women 12–18 years old (-1.05 ; 95% CI -1.32 to -0.78 ; $p < 0.0001$; $I^2 = 0\%$) and in women 19–35 years (-1.04 ; 95% CI -1.86 to -0.21 ; $p = 0.001$, $I^2 = 30\%$). However, in women 36–49 years old, there

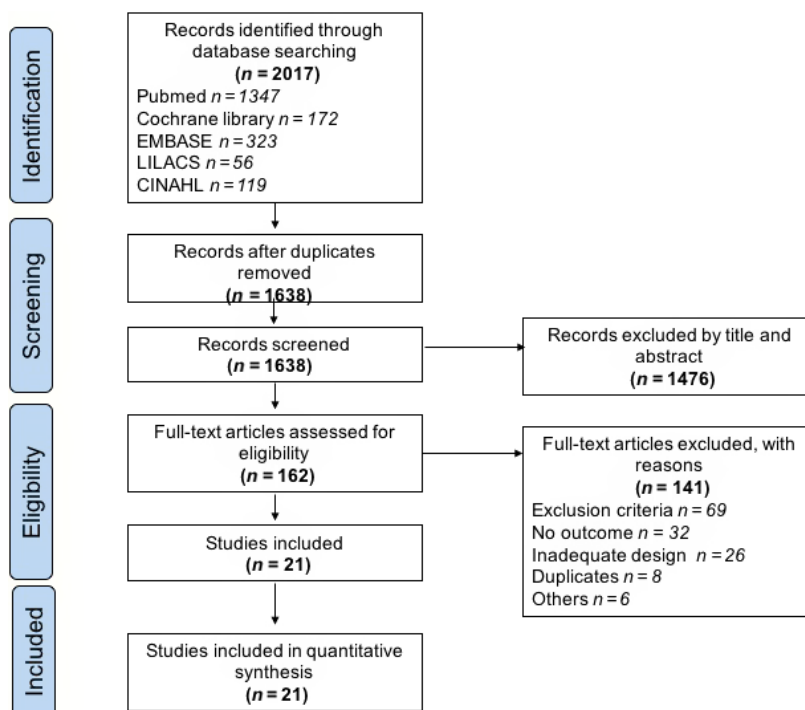


Figure 1. Flow chart of screening and selection of studies.

Table 1. Characteristics of included studies.

First author, year (ref. no.)	Country	Type of study	Population characteristic	Database analyzed	Age range or mean \pm SD	Sample size	Reported outcome	SR ^a method	Time lag between SR ^a and DM ^b
Brettschneider, 2011 (16)	Germany	Cross-sectional	General population	KiGGS ^g	14–17	948	MW,MH, MBMI, DW,DH, DBMI	Personal interview	Same day
Brunner, 2007 (17)	USA	Cross-sectional	General population	CHIC ⁱ	18–25 26–35 36–49	89 138 48	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	Same day
Carvalho, 2014 (18)	Brazil	Cross-sectional	General population	ISA-Capital ^h	12–19	32	MW,MH, MBMI	Long-distance survey (telephone)	Different days (non-specified)
Ekstrom, 2015 (19)	Stockholm	Prospective cohort	General population	–	16.5 + 0.4	889	MW,MH, MBMI, DW,DH, DBMI	Long-distance survey (online)	NR
Elgar, 2005 (20)	Wales	Cross-sectional	High school students	HBSC Study ^f	15–17	211	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	Same day
Fonseca, 2009 (21)	Portugal	Cross-sectional	High school students	HBSC Study ^f	14 + 1.8	233	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	Same day
Galán, 2001 (22)	Spain	Cross-sectional	High school students	–	15–18	1810	DW,DH, DBMI	Self-administered survey	Same day
Himes, 2001 (23)	USA	Cross sectional	General population	NHANES III ^e	12–16	876	MW,MH, MBMI, DW,DH, DBMI	NR	Same day
Larsen, 2007 (24)	Netherlands	Cross-sectional	University students	–	20.9 + 2.40	209	MW,MH, MBMI	Questionnaire on-site	Same day
Leatherdale, 2013 (25)	Canada	Cross-sectional	High school students	–	14–15	65	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	Different days (1 week)
Legleye, 2014 (26)	France	Cross-sectional	General population	ESCAPAD ^d	17–18	140	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	Same day
Lo, 2011 (27)	China	Cross-sectional	High school students	HKSOS project ^c	13.67 + 1.18 16.29 + 0.98	1838 1275	MW,MH, MBMI	Questionnaire on-site	NR
Marrodan, 2013 (28)	Spain	Cross-sectional	General population	–	18–24 25–34 35–44	181 1486 1876	MW,MH, MBMI	Personal interview	Same day
Peixoto, 2006 (29)	Brazil	Cross-sectional	General population	–	20–24 25–34 35–44	97 184 150	DW,DH, DBMI	Personal interview	Same day
Pregolato, 2009 (30)	Brazil	Cross-sectional	University students	–	28.3 + 11	549	MW,MH, MBMI	NR	Same day
Pursey, 2014 (31)	Australia	Cross sectional	General population	–	18–35	93	MW, MH	Long-distance survey (online)	Different days (<1 month)

Table 1. Continued

First author, year (ref. no.)	Country	Type of study	Population characteristic	Database analyzed	Age range or mean \pm SD	Sample size	Reported outcome	SR ^a method	Time lag between SR ^a and DM ^b
Rodrigues, 2013 (32)	Brazil	Cross-sectional	High school students	–	14–19	40	MW,MH, MBMI, DW,DH, DBMI	NR	Same day
Savane, 2013 (33)	Spain	Prospective cohort	University students	–	18–37	476	MW,MH, MBMI, DW,DH, DBMI	Self-administered survey	NR
Shin, 2014 (6)	USA	Cross-sectional	General population	NHANES ^j	16–25 26–35 36–44	1252 592 599	MW	Self-administered survey	Same day
Unikel Santocini, 2009 (35)	Mexico	Cross-sectional	High school students	–	15–19	2357	MW,MH, DW,DH	Self-administered survey	Same day
Vitale 2013 (36)	Argentina	Cross-sectional	High school students	–	15–18	61	MW,MH, MBMI	Self-administered survey	Same day

DBMI, difference body mass index; DH, difference height; DM, direct measured; DW, difference weight; MBMI, measure body mass index; MH, measure height; MW, measure weight; NR, not reported; SR, self-reported.

^aSelf-reported.

^bDirectly measured.

^cHong Kong Student Obesity Surveillance (HKSOS) project.

^dESCAPAD survey (Survey on health and behavior).

^eNational Health and Nutrition Examination Survey III.

^fHealth Behavior School-Aged Children (HBSC).

^gGerman Health Interview and Examination Survey for Children and Adolescents (KiGGS).

^hHealth Survey of São Paulo (ISA-Capital).

ⁱThe Contraceptive History, Initiation, and Choice (CHIC) Study.

^jNational Health and Nutrition Examination.

was no statistically significant difference between methods (-0.26 ; 95% CI -0.99 to 0.44]; $p = 0.49$; $I^2 = 0\%$) (Figure 2a).

The results by region were in the same direction in all three meta-analyzed regions (Figure 3a). The difference between self-reported and directly measured weight was -1.14 kg (95% CI -1.67 to -0.61 ; $p < 0.0001$; $I^2 = 0\%$) in Latin America and the Caribbean, -1.02 kg (95% CI -1.68 to -0.37]; $p = 0.002$, $I^2 = 55\%$) in Europe and -1.51 kg (95% CI -2.53 to -0.48 ; $p = 0.004$; $I^2 = 0\%$) in North America. We only found one study for Asia and one for Oceania, and they were not included in the meta-analysis (24,28).

In the analysis by time of data collection we found that if obtained within the same day there was a -0.97 kg difference (95% CI -1.37 to -0.57 ; $p < 0.001$; $I^2 = 15\%$) between self-reported and directly measured weight. No statistically significant difference was found when data were collected on separate days (-1.64 kg; 95% CI -4.30 to 1.03 ; $p = 0.23$; $I^2 = 0\%$) (Figure 4a).

We also evaluated the influence of the self-reported method used when compared with directly measured data (Figure S1A). The analysis suggested that there was a negative difference if the information was gathered through a long-distance survey (-1.46 kg; 95% CI -2.27 to -0.64 ; $p = 0.0004$; $I^2 = 0\%$) or a self-administered questionnaire on-site (-1.14 kg; 95% CI -1.79 to -0.48 ; $p = 0.006$; $I^2 = 54\%$). The difference was lower when gathered during an in-person interview (-0.27 kg; 95% CI -0.80 to 0.25 ; $p = 0.74$; $I^2 = 46\%$).

Only two studies classified their population according to BMI status of participants. We found that those who were overweight underestimated their weight by -0.39 kg (95% CI -0.59 to -0.19 ; $p = 0.0001$; $I^2 = 0\%$) (28,29). Because of the high heterogeneity between the studies, for the other three BMI categories (underweight, normal weight or obesity) we found no statistically significant results (Figure 5a).

As mentioned previously, two studies only reported mean difference between methods, without specifying

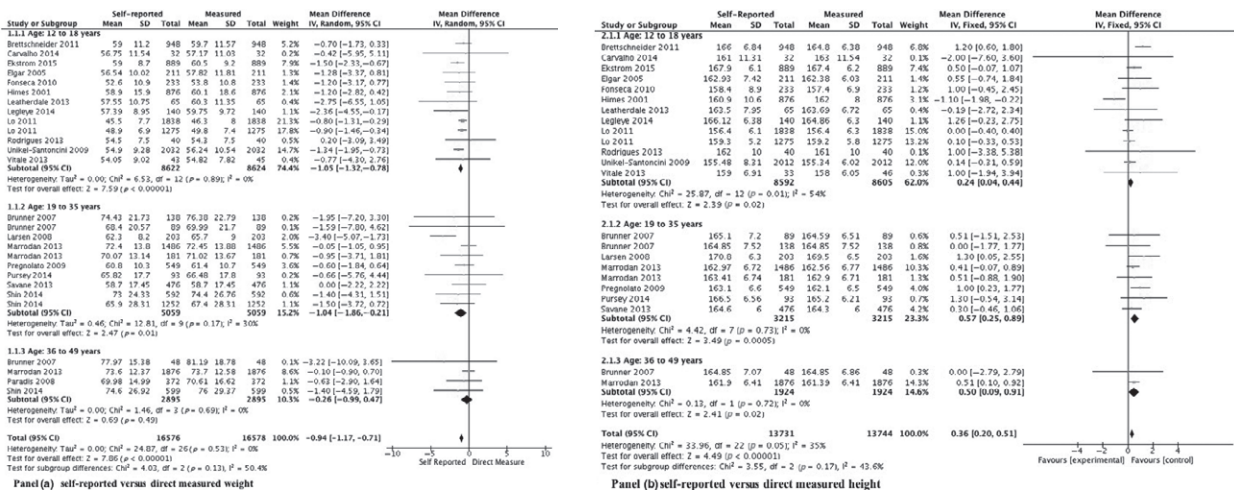
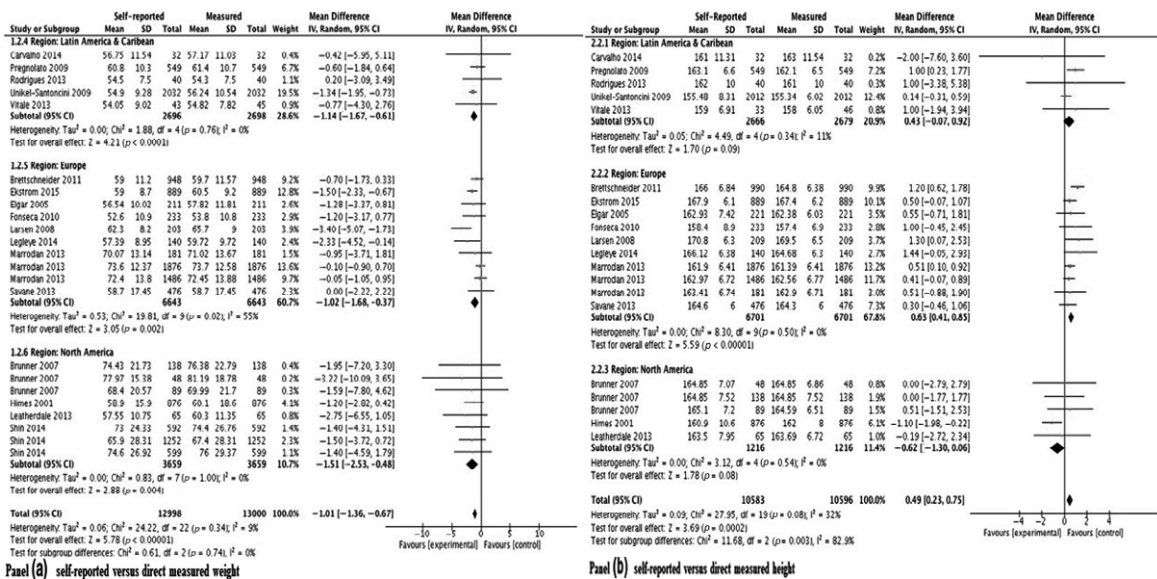


Figure 2. Forest plot of self-reported vs. direct measured weight (a) and height (b) (mean difference) by age group.



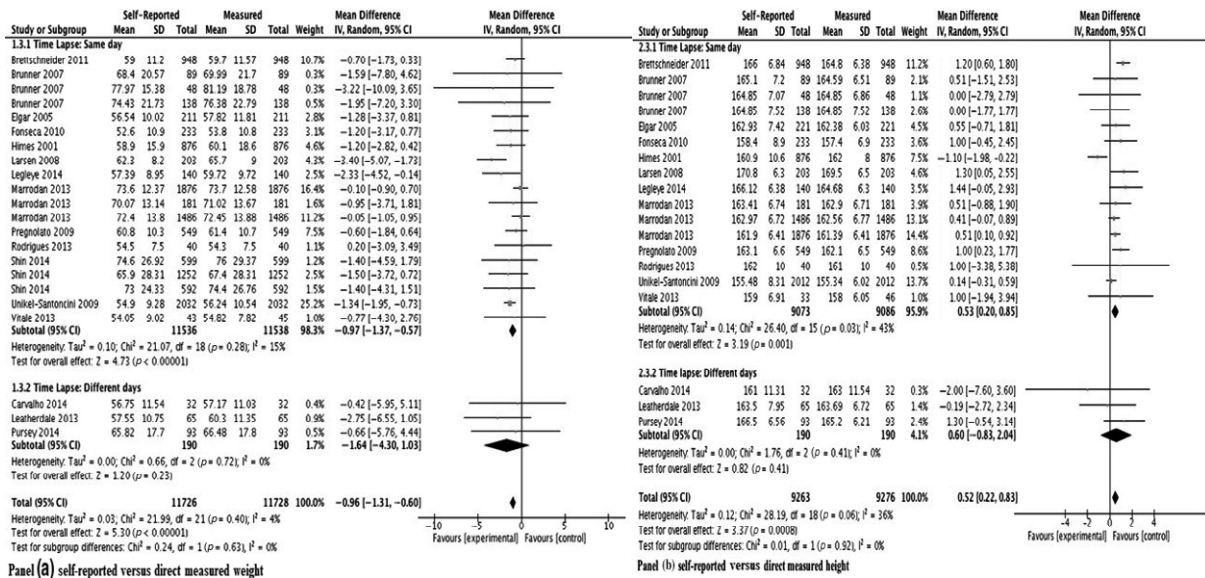


Figure 4. Forest plot of self-reported vs. direct measured weight (a) and height (b) (mean difference) by time between self-reported and direct measurements.

$I^2 = 54%$); 19–35 years, mean difference = 0.57 cm (95% CI 0.25–0.89; $p < 0.001$; $I^2 = 0%$); 36–49 years, mean difference = 0.50 cm (95% CI 0.09–0.91; $p = 0.02$; $I^2 = 0%$).

Analysis by region showed a significant mean difference between self-reported height and directly measured height of 0.63 cm (95% CI 0.41–0.85; $p < 0.0001$; $I^2 = 0%$) in Europe. No statistical differences were found in the Americas: North America: -0.62 cm (95% CI -1.30 to

0.06; $p = 0.08$; $I^2 = 0%$); Latin America and the Caribbean: 0.43 cm (95% CI -0.07 to 0.92; $p = 0.09$; $I^2 = 11%$) (Figure 3b). Two studies were excluded from the meta-analysis because each represented the only reference from their region (27,31).

In the analysis by time of data collection we found that if obtained within the same day, the difference was 0.53 cm (95% CI 0.20–0.85; $p = 0.001$; $I^2 = 43%$). No significant difference was found when obtained on separate days

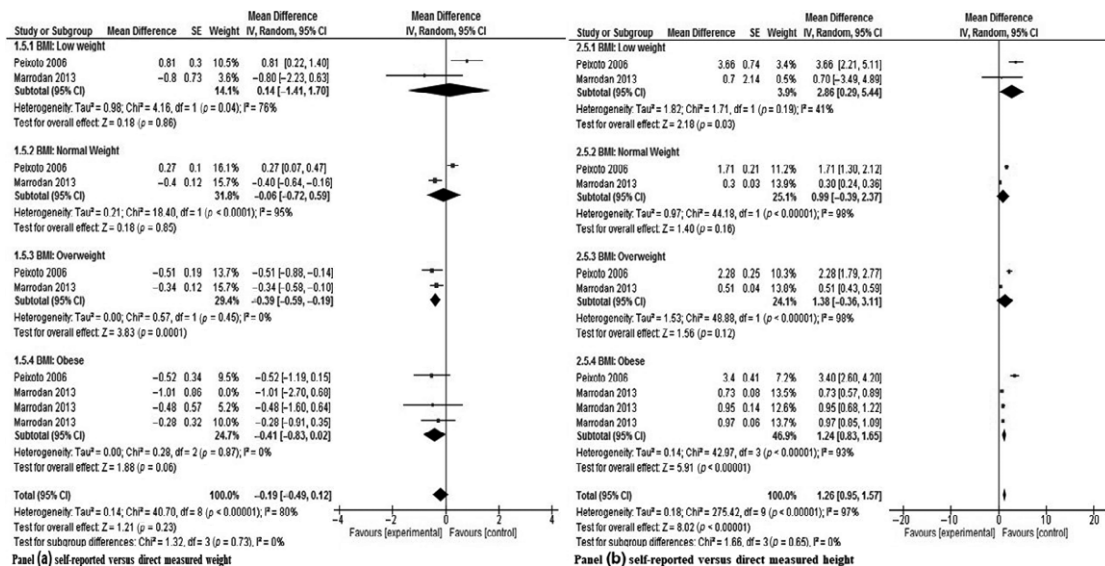


Figure 5. Forest plot of self-reported vs. direct measured weight (a) and height (b) (mean difference) by body mass index (BMI).

(0.60 cm; 95% CI -0.83 to 2.04 ; $p = 0.41$; $I^2 = 0\%$) (Figure 4b).

We also evaluated the influence of the specific self-reported method used when compared with directly measured height data (Figure S1B). The analysis showed a significant difference if the information was gathered through a long-distance survey (0.55 cm; 95% CI 0.00 – 1.09 ; $p = 0.05$; $I^2 = 0\%$) or in an in-person interview (0.65 cm; 95% CI 0.28 – 1.02 ; $p = 0.0005$; $I^2 = 38\%$). No statistically significant difference was found when the data were gathered through an on-site self-administered questionnaire (0.10 cm; 95% CI -0.68 to 0.47 ; $p = 0.72$; $I^2 = 70\%$).

The high heterogeneity found between studies in the subgroup analysis based on women's BMI categories prevented us from obtaining an estimate difference between self-reported and directly measured height (Figure 5b) (28,29).

The separate analysis for the two studies (27,31) reporting only the mean difference between methods found that the results were consistent with the findings previously presented, and showing self-reported height greater than direct measurements. There was no heterogeneity between studies ($I^2 = 0\%$) (Figure S2B).

Discussion

The results of this review showed an overall underestimation of weight (-0.94 kg) and an overestimation of height ($+0.36$ cm) when comparing self-reported with directly measured values in women of reproductive age.

In the prespecified subgroup analyses, the findings remained consistent. We found that women aged 12–35 years underreported their weight by 0.78–1.17 kg. Older women also underreported their weight but the difference was not statistically significant in this age group. The underestimation of self-reported weight was found in all studied regions, reaching a mean difference between self-reported and direct measured weight as high as 1.50 kg in North America. Few studies presented data in overweight women; the results on weight were similar to normal weight women. It was not possible to estimate the differences in underweight or obese subgroups.

We found that the underestimation of weight persisted if data were collected via an on-site self-administered questionnaire or a long-distance survey (online or via telephone); however, when self-reported data were collected by on-site in-person interviews, this underestimation was lower and was not statistically significant.

Regarding height, the results showed a consistent overestimation in all age groups. These findings were also observed in studies from Europe and North America but not in those from Latin America or the Caribbean. The

overestimation of height persisted when collected with an on-site in-person interview or long-distance survey; however, there was no statistically significant difference with directly measured values when using an on-site self-administered questionnaire. Our results confirmed the data published by Gorber et al. (4) in the general population and updated by Engstrom et al.'s (36) results from 2002 to 2015. Both these studies showed an underestimation of weight and overestimation of height. In our study, as well as that of Gorber et al.'s and Engstrom et al.'s reviews, the standard deviations were large in all included studies, suggesting a significant variability between women concerning the accuracy of self-reported height and weight measurements.

One of the authors carried out three pilot tests of search strategies on MEDLINE to explore the potential sensitivity and specificity of the electronic searches. We assume that the risk of publication bias is low (Table S1). Poor reporting of studies was the major problem found when assessing the risk of bias of included studies. To address this limitation, we contacted the primary authors of those articles with missing data.

Although large numbers of women have been studied, Asia and Oceania had little representation in the final selection of studies, with only one article from each region. Moreover, some of the included studies had a relatively small sample size.

One limitation of our review was the high heterogeneity found when the meta-analysis combined studies reporting means and those reporting only mean differences. To compensate for this limitation, we presented a separate meta-analysis for those studies reporting only a mean difference. The main strength of this review is that, by restricting the population inclusion criteria, we could control for the large heterogeneity between studies and calculate a reliable summary estimate that quantifies the bias that occurs when using self-reported weight and height data for women of reproductive age.

Finally, we observed that there is a difference in relation to the degree of significance in some analyses. In this regard, the limited number of studies in some sub-analyses challenged the interpretation of the results.

This review presents the difference of using self-reported weight and height compared with direct measurements in women of reproductive age with no eating disorders or conditions that might confound the comparison. The population selected in this study allowed us to reduce the heterogeneity between studies and to achieve a summary estimate of possible bias.

Self-reported maternal weight and height are widely used, particularly in situations where even basic anthropometric measurements cannot be taken. Self-reported measures are used in clinical practice and in studies that

relate them to pregnancy outcomes. This review shows a low bias in the estimation of weight and height using self-reported measures; for example, the BMI of a woman with a weight of 50 kg and a height of 1.65 m, would differ by 2.36% (95% CI 2.07–2.58%) if measured using self-reported data. Our interpretation is that self-reported weight and height in women of reproductive age is a measure that closely estimates the real values and can be used as proxy of real values both in clinical and research evaluation.

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References

1. WHO. Obesity: preventing and managing the global epidemic. Report of a WHO Consultation. Geneva: WHO, 2000.
2. McAdams MA, Van Dam RM, Hu FB. Comparison of self-reported and measured BMI as correlates of disease markers in US adults. *Obesity*. 2007;15:188–96.
3. Hanson MA, Bardsley A, De-Regil LM, Moore SE, Oken E, Poston L, et al. The International Federation of Gynecology and Obstetrics (FIGO) recommendations on adolescent, preconception, and maternal nutrition: “Think Nutrition First”. *Int J Gynaecol Obstet*. 2015;131(Suppl 4): S213–53.
4. Connor Gorber S, Tremblay M, Moher D, Gorber B. A comparison of direct vs. self-report measures for assessing height, weight and body mass index: a systematic review. *Obesity Rev*. 2007;8:307–26.
5. Weaver EH, Gibbons L, Belizan JM, Althabe F. The increasing trend in preterm birth in public hospitals in northern Argentina. *Int J Gynaecol Obstet*. 2015;130:137–41.
6. Shin D, Chung H, Weatherspoon L, Song WO. Validity of prepregnancy weight status estimated from self-reported height and weight. *Matern Child Health J*. 2014;18:1667–74.
7. Rubinstein F, Micone P, Bonotti A, Wainer V, Schwarcz A, Augustovski F, et al. Influenza A/H1N1 MF59 adjuvanted vaccine in pregnant women and adverse perinatal outcomes: multicentre study. *BMJ*. 2013;346:f393.
8. Acevedo P, Lopez-Ejeda N, Alferez-Garcia I, Martinez-Alvarez JR, Villarino A, Cabanas MD, et al. Body mass index through self-reported data and body image perception in Spanish adults attending dietary consultation. *Nutrition*. 2014;30:679–84.
9. Moher D, Liberati A, Tetzlaff J, Altman DG, PRISMA Group. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *Ann Intern Med*. 2009;151:264–9, W64.
10. Liberati A, Altman DG, Tetzlaff J, Mulrow C, Gotzsche PC, Ioannidis JP, et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate healthcare interventions: explanation and elaboration. *BMJ*. 2009;339:b2700.
11. Bardach A, Ciapponi A, Garcia-Marti S, Glujovsky D, Mazzoni A, Fayad A, et al. Epidemiology of acute otitis media in children of Latin America and the Caribbean: A systematic review and meta-analysis. *Int J Pediatr Otorhinolaryngol*. 2011;75:1062–70.
12. Vandembroucke JP, von Elm E, Altman DG, Gotzsche PC, Mulrow CD, Pocock SJ, et al. Strengthening the reporting of observational studies in epidemiology (STROBE): explanation and elaboration. *PLoS Med*. 2007;4:e297.
13. Review Manager (RevMan) 5.2 edn. Copenhagen: The Nordic Cochrane Centre, 2012.
14. Landsberg B, Bastian I, Plachta-Danielzik S, Lange D, Johannsen M, Seiberl J, et al. Self-reported and measured height and weight in adolescents. *Gesundheitswesen*. 2011;73:40–5.
15. Avila-Funes JA, Gutiérrez-Robledo LM, De Leon Ponce, Rosales S. Validity of height and weight self-report in Mexican adults: results from the national health and aging study. *J Nutr Health Aging*. 2004;8:355–61.
16. Brettschneider AK, Rosario AS, Ellert U. Validity and predictors of BMI derived from self-reported height and weight among 11- to 17-year-old German adolescents from the KiGGS study. *BMC Res Notes*. 2011;4:414.
17. Brunner Huber LR. Validity of self-reported height and weight in women of reproductive age. *Matern Child Health J*. 2007;11:137–44.
18. Carvalho AM, Piovezan LG, Selem SS, Fisberg RM, Marchioni DM. Validation and calibration of self-reported weight and height from individuals in the city of Sao Paulo. *Rev Bras Epidemiol*. 2014;17:735–46.
19. Ekstrom S, Kull I, Nilsson S, Bergstrom A. Web-based self-reported height, weight, and body mass index among Swedish adolescents: a validation study. *J Med Internet Res*. 2015;17:e73.
20. Elgar FJ, Roberts C, Tudor-Smith C, Moore L. Validity of self-reported height and weight and predictors of bias in adolescents. *J Adolesc Health*. 2005;37:371–5.
21. Fonseca H, Silva AM, Matos MG, Esteves I, Costa P, Guerra A, et al. Validity of BMI based on self-reported weight and height in adolescents. *Acta Paediatr*. 2010;99:83–8.
22. Galan I, Gandarillas A, Febrel C, Meseguer C. [Validation of self-reported weight and height in an adolescent population]. *Gaceta sanitaria/SESPAS*. 2001;15:490–7.
23. Himes JH, Faricy A. Validity and reliability of self-reported stature and weight of US adolescents. *Am J Hum Biol*. 2001;13:255–60.

24. Larsen JK, Ouwens M, Engels RC, Eisinga R, van Strien T. Validity of self-reported weight and height and predictors of weight bias in female college students. *Appetite*. 2008;50:386–9.
25. Leatherdale ST, Laxer RE. Reliability and validity of the weight status and dietary intake measures in the COMPASS questionnaire: are the self-reported measures of body mass index (BMI) and Canada's food guide servings robust? *Int J Behav Nutr Phys Act*. 2013;10:42.
26. Legleye S, Beck F, Spilka S, Chau N. Correction of body-mass index using body-shape perception and socioeconomic status in adolescent self-report surveys. *PLoS ONE*. 2014;9:e96768.
27. Lo WS, Ho SY, Wong BY, Mak KK, Lam TH. Validity and test-retest reliability in assessing current body size with figure drawings in Chinese adolescents. *Int J Pediatr Obes*. 2011;6:e107–13.
28. Marrodan MD, Martinez-Alvarez JR, Villarino A, Alferrez-Garcia I, Gonzalez-Montero de Espinosa M, Lopez-Ejeda N, et al. Utilidad de los datos antropométricos auto-declarados para la evaluación de la obesidad en la población española; estudio EPINUT-ARKOPHARMA. [Utility of self-reported anthropometric data for evaluation of obesity in the Spanish population; study EPINUT-ARKOPHARMA]. Article in Spanish. *Nutr Hosp*. 2013;28:676–82.
29. Peixoto MdRG, D'Aquino Benicio MHD, Veiga Jardim PC. Validity of self-reported weight and height: the Goiânia study, Brazil. *Rev Saude Publica* 2006;40:1065–72.
30. Pregolato TdS, Mesquita LM, Ferreira PG, Santos MM, Santos CC et al. Validade de medidas autorreferidas de massa e estatura e seu impacto na estimativa do estado nutricional pelo índice de massa corporal. [Validity of self-reported measures of mass and height and its impact on the estimation of nutritional status by body mass index.] In Portuguese. *Rev Bras Crescimento Desenvolv Hum*. 2009;19:35–41.
31. Pursey K, Burrows TL, Stanwell P, Collins CE. How accurate is web-based self-reported height, weight, and body mass index in young adults? *J Med Internet Res*. 2014;16:e4.
32. Rodrigues PRM, Gonçalves-Silva RMV, Pereira RA. Validity of self-reported weight and stature in adolescents from Cuiabá. Central-Western Brazil. *Rev Nutr*. 2013;26:283–90.
33. Savane FR, Navarrete-Munoz EM, Garcia de la Hera M, Gimenez-Monzo D, Gonzalez-Palacios S, Valera-Gran D, et al. Validez del peso y talla auto-referido en población universitaria y factores asociados a las discrepancias entre valores declarados y medidos. [Validation of self-reported weight and height university population and factors associated with differences between self reported and measured antropometrics]. In Spanish. *Nutr Hosp*. 2013;28:1633–8.
34. Unikel-Santoncini C, Ocampo-Ortega R, Zambrano-Ruiz J. Exactitud del autorreporte de peso y talla en mujeres de 15 a 19 años del Estado de México. [Exactitude of self-reported weight and height in 15 to 19 year old female adolescents of the State of Mexico]. In Spanish. *Salud Publica Mex*. 2009;51:194–201.
35. Vitale R, Lavin Fueyo J, Rivera C, Mamondi V, Berra S. Validez del peso y la talla declarados en adolescentes escolarizados de la ciudad de Córdoba. [Validity of self-reported weight and height by high school students in the city of Cordoba city]. *Rev Salud Pública (Córdoba)*. 2013;17:42–8.
36. Engstrom JL, Paterson SA, Doherty A, Trabulsi M, Speer KL. Accuracy of self-reported height and weight in women: an integrative review of the literature. *J Midwifery Womens Health*. 2003;48:338–45.

Supporting information

Additional Supporting Information may be found in the online version of this article:

Figure S1. Forest plot of self-reported vs. direct measured weight (A) and height (B) (mean difference) by self-report method.

Figure S2. Forest plot of mean difference between self-reported and direct measured weight (A) and height (B) in studies that only reported mean differences.

Table S1. Assessment of risk of bias by article.

Appendix S1. MEDLINE search strategy.